

Recombination fraction in pre-recombinant inbred lines (PRERIL) - revisiting a century old problem in genetics

Shizhong Xu^{1*} and José Osorio y Fortéa²

Abstract

Background Traditional recombinant inbred lines (RILs) are generated from repeated self-fertilization or brothersister mating from the F_1 hybrid of two inbred parents. Compared with the F_2 population, RILs cumulate more crossovers between loci and thus increase the number of recombinants, resulting in an increased resolution of genetic mapping. Since they are inbred to the isogenic stage, another consequence of the heterozygosity reduction is the increased genetic variance and thus the increased power of QTL detection. Self-fertilization is the primary form of developing RILs in plants. Brother-sister mating is another way to develop RILs but in small laboratory animals. To ensure that the RILs have at least 98% of homozygosity, we need about seven generations of self-fertilization or 20 generations of brother-sister mating. Prior to homozygosity, these lines are called pre-recombinant inbred lines (PRERIL). Phenotypic values of traits in PRERILs are often collected but not used in QTL mapping. To perform QTL mapping in PRERILs, we need the recombination fraction between two markers at generation *t* for *t*<7 (selfng) or *t*<20 (brother-sister mating) so that the genotypes of QTL fanked by the markers can be inferred.

Results In this study, we developed formulas to calculate the recombination fractions of PRERILs at generation *t* in self-fertilization, brother-sister mating, and random mating. In contrast to existing works in this topic, we used computer code to construct the transition matrix to form the Markov chain of genotype array between consecutive generations, the so-called recurrent equations.

Conclusions We provide R functions to calculate the recombination fraction using the newly developed recurrent equations of ordered genotype array. With the recurrent equations and the R code, users can perform QTL mapping in PRERILs. Substantial time and effort can be saved compared with QTL mapping in RILs.

Keywords Computer generated transition matrix, Markov chains, Recombinant inbred lines, Recurrent equations

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Background

Herbert Spencer Jennings (1868–1947), an American Professor, was the frst geneticist to study the behaviors of recombinant inbred lines (RILs) [\[9](#page-17-0)]. At that time, recombinant inbred lines had not been conceptualized. Jennings [\[8](#page-17-1), [9\]](#page-17-0) called the two copies of a single Mendelian locus a pair of Mendelian characters while called alleles from two loci two pairs of Mendelian characters, where Mendelian factors are referred to as genetic units passed from parents to ofspring. Jennings described the process of generating RILs as repeated self-fertilization,

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Jennings' [\[8](#page-17-1), [9\]](#page-17-0) work was fundamental but very disorganized in its written form, unfortunately. It is not until Robbins [[16\]](#page-17-3) who redescribed Jennings' work in a clear and well organized manner, that Jennings' [[9\]](#page-17-0) work became well-known. Two mating systems (random mating and self-fertilization) introduced by Jennings and reintroduced by Robbins will be discussed in this study. However, we mainly cited the work by Robbins [[14](#page-17-4)[–16](#page-17-3)]. Both Jennings and Robbins defned the parameter of interest as linkage ratio denoted by *r*. But their *r* and the *r* in modern genetics are entirely different. The *r* defined as linkage ratio is a relative number of the parental types of gametes compared with the recombinant types of gametes $[9]$ $[9]$ $[9]$. The *r* defined as the recombination fraction in modern genetics is a proportion of the recombinant gametes over all possible gametes in a population of interest. To avoid any potential confusion, we denote the recombination fraction in modern genetics by θ to avoid using *r* as the recombination fraction.

Robbins' random mating recurrent equations are clearer. His equations lead to the proportions of the four types of gametes expressed as functions of *r* and the number of generations of random mating. He concluded that (1) in random mating, the efect of incomplete linkage between two factors is only temporary and (2) continued random mating results in a population in which the distribution of "B" factors among the "A" and "a" factors is the same as the distribution of the "b" factors among the "A" and "a" factors. The second conclusion is simply the statistical independence between the two factors or linkage equilibrium after many generations of random mating. In fact, the two conclusions imply the same consequence in modern quantitative genetics: genetic correlation caused by incomplete linkage is temporary [[12](#page-17-5)].

Robbins [[16\]](#page-17-3) reinvestigated all problems raised by Jennings [[9\]](#page-17-0) and extended the work into a higher level. Especially for the random mating system where he, after extensive derivation, developed a functional relationship of the gametic frequencies to the initial gametic frequencies using the sum of geometric series. As demonstrated in Supplementary Note S3, the functional relationship of the recombination fraction of Robbins is identical to the functional relationship developed by Darvasi and Soller [[4](#page-17-6)] who used an extremely simple method. Darvasi and Soller [[4\]](#page-17-6) called the lines generated from such a repeated random mating scheme the advanced intercross lines (AIL). They frst derived the recurrent equation for the recombination fraction starting with the F_1 hybrid of a cross between two inbred lines. From the recurrent equation, they expressed the recombination fraction at generation *t* as a function of the recombination fraction in the original population,

$$
\theta_t = \frac{1}{2} \Big[1 - (1 - 2\theta)(1 - \theta)^{t-2} \Big] \tag{1}
$$

where $t \geq 2$ with $t = 2$ being the F_2 population.

Robbins' [[16](#page-17-3)] other contribution to the subject was to reinvestigate the recurrent genotypic frequencies in the self-fertilization system. He pooled the $4 \times 4 = 16$ total genotypes with phase information into 10 distinguished unphased genotypes. The recurrent equations were much cleaner than those given by Jennings [\[9](#page-17-0)], although Robbins inherited Jennings's notation system, e.g., using the same *r* to represent the linkage ratio and the same *p*, *q*, *s*, *t* notation to denote the four gametic frequencies in the random mating system. Robbins did not provide the asymptotic recombination fraction after infnite number of generations with self-fertilization.

Haldane and Waddington [\[6](#page-17-7)] developed the recombination fractions at the equilibrium stage after infnite number of self-fertilization and brother-sister mating. Haldane and Waddington [[6\]](#page-17-7) combined some of the 10 unphased genotypes proposed by Robins [\[16](#page-17-3)] into a common class and yielded 5 composite genotypes. They delt with only 5 recurrent equations, significantly reduced the computational complexity. The major contribution of the Haldane and Waddington's study [\[6](#page-17-7)] was the brother-sister matting system for linkage analysis, which was not touched by previous authors for two pairs of linked characters. Haldane and Waddington [\[6](#page-17-7)] developed the $10 \times 10 = 100$ recurrent equations for the genotypes of the sibling pairs. The absorption of the original $16 \times 16 = 256$ fully phased recurrent equations into the $10 \times 10 = 100$ unphased recurrent equations represents a substantial reduction in computation. The recurrent equations convert the frequencies of the 100 genotype combinations from the previous generation to the genotype frequencies of the current generation using a 10×10 transition matrix in the Markov chain system.

Other than the recurrent equation of the recombination fraction developed by Darvasi and Soller [\[4\]](#page-17-6), none of the previous works reported the recombination fraction before the consecutive mating systems reach equilibrium. The recurrent equations for genotype frequencies under self-fertilization and brother-sister mating were all derived manually, which are prone to error when a computer program code is written. Broman [\[2\]](#page-17-8) extended the asymptotic recombination fraction of RILs of brothersister mating from an 8-way crosses and showed that the fnal recombination fraction is

$$
\rho_{SW} = \frac{7\theta}{1 + 6\theta} \tag{2}
$$

No recurrent equations are provided to determine the recombination fraction before the lines reach the equilibrium value. The purposes of this study are to present (1) a derivation of the recombination fraction at generation $(t < \infty)$ before the system reaches the equilibrium, (2) a computer code generated transition matrix for recurrent equations of genotype frequencies. Relevant background knowledge and recombination fraction at generation $(t < \infty)$ from works of previous authors are given in the Supplementary Note S1, Note S2 and Note S3.

Methods

Basic defnition

Consider two linked loci (A and B) with a recombination fraction of θ for $0 < \theta < 0.5$. Define the diploid genotypes of the two inbred parents that initiate the F_1 hybrid by *AB/AB* and *ab/ab*, respectively. The genotype of the F_1 hybrid is AB/ab . In each genotype, the maternal and paternal gametes are separated by a slash. The four possible gametes from this F_1 hybrid are *AB*, *Ab*, *aB* and *ab* with probabilities $\frac{1}{2}(1-\theta)$, $\frac{1}{2}\theta$, $\frac{1}{2}\theta$ and $\frac{1}{2}(1-\theta)$, respectively. The gametes of the F_1 hybrid make the genotypes of the F_2 population. Therefore, the recombination fraction of the F_2 generation is

The mating system will start with $t=1$, i.e., the F_1 generation, from which the recurrent equations of genotypes will be developed. The $4\times4=16$ possible genotypes in the $F₂$ population are shown in Table [1](#page-2-0) below.

In the current literature, the recombination fraction is often denoted by *r*. However, Jennings [\[9](#page-17-0)] and Robbins (1917) defned *r* as a linkage ratio parameter with an entirely different interpretation. They set each of the recombinant gametes to 1, and each of the parental gametes to r relative to the recombinant gamete. The relative contributions of the four gametes are r from AB or ab, and 1 from Ab or aB. To avoid notational confusion, we denote the recombination fraction by θ . The relationship between θ and r is

$$
\theta = \frac{1}{1+r} \text{ or } r = \frac{1-\theta}{\theta} \tag{4}
$$

Starting from Table [1,](#page-2-0) the recurrent equations of genotype and gamete frequencies are developed for self-fertilization, brother-sister mating, and random mating. These recurrent equations allow us to calculate the recombination fractions of PRERILs at generation *t*. We start with self-fertilization, then brother-sister mating, and fnally random mating. We assume that the recombination fractions are the same for the male and female gametes. Haldane and Waddington [[6\]](#page-17-7) denoted the recombination fraction for the female gamete by β and for the male gamete by δ . They intended to cover insects, which often have diferent recombination fractions between sexes. We do not diferentiate the male and female gametes and thus the results of this study only apply to diploid plants and diploid animals where $\beta = \delta = \theta$ is the recombination fraction.

The 16 fully phased genotypes in Table [1](#page-2-0) are rearranged into a column vector with the order shown in Table [2](#page-3-0). Verbally, the four genotypes of the frst row in Table [1](#page-2-0) become the first four genotypes in the 16×1 vector of Table [2](#page-3-0). Gametic probabilities that each of the

$$
\theta_2 = \frac{\Pr\left(Ab\right) + \Pr\left(ab\right)}{\Pr\left(AB\right) + \Pr\left(ab\right) + \Pr\left(Ab\right) + \Pr\left(ab\right)} = \frac{\frac{1}{2}\theta + \frac{1}{2}\theta}{\frac{1}{2}(1-\theta) + \frac{1}{2}\theta + \frac{1}{2}\theta + \frac{1}{2}(1-\theta)} = \theta
$$
\n(3)

Table 1 The 16 possible genotypes of the F_2 population from the F_1 hybrid with genotype *AB/ab*

Female\Male	$AB_{\overline{2}}^{\perp}(1-\theta)$	$Ab^{\perp}_{\overline{2}}\theta$	$aB\frac{1}{2}\theta$	$ab_{\overline{2}}(1-\theta)$
$AB^1_2(1-\theta)$	$AB/AB\frac{1}{4}(1-\theta)^2$	$AB/Ab\frac{1}{4}\theta(1-\theta)$	$AB/AB\frac{1}{4}\theta(1-\theta)$	$AB/ab\frac{1}{4}(1-\theta)^2$
$Ab_{\overline{2}}^{1}\theta$	$Ab/AB\frac{1}{4}\theta(1-\theta)$	Ab/Ab $\frac{1}{4}\theta^2$	Ab/aB $\frac{1}{4}\theta^2$	$Ab/ab \frac{1}{4}\theta(1-\theta)$
$aB_2^1\theta$	$aB/AB\frac{1}{4}\theta(1-\theta)$	a B/Ab $\frac{1}{4}\theta^2$	$aB/aB\frac{1}{4}\theta^2$	$aB/ab\frac{1}{4}\theta(1-\theta)$
$ab_{\overline{2}}(1-\theta)$	$ab/AB\frac{1}{4}(1-\theta)^2$	$ab/Ab\frac{1}{4}\theta(1-\theta)$	$ab/ab\frac{1}{4}\theta(1-\theta)$	$ab/ab \frac{1}{4}(1-\theta)^2$

Xu and Osorio y Fortéa *BMC Genomics (2024) 25:822* Page 4 of 18

Table 2 Gametic probability table (the *H* matrix) from each of the 16 fully phased genotypes

Entry	Genotype	ΑB	Ab	аB	ab
1	AB/AB	1	Ω	0	Ω
2	AB/Ab	1/2	1/2	0	Ω
3	AB/AB	1/2	0	1/2	0
4	AB/ab	$rac{1}{2}(1-\theta)$	$rac{1}{2}\theta$	$rac{1}{2}\theta$	$rac{1}{2}(1-\theta)$
5	Ab/AB	1/2	1/2	0	0
6	Ab/Ab	0	1	0	0
7	Ab/ab	$rac{1}{2}\theta$		$\frac{1}{2}(1-\theta) = \frac{1}{2}(1-\theta)$	$rac{1}{2}\theta$
8	Ab/ab	0	1/2	$\mathbf 0$	1/2
9	aB/AB	1/2	0	1/2	0
10	aB/Ab	$rac{1}{2}\theta$	$rac{1}{2}(1-\theta)$	$rac{1}{2}(1-\theta)$	$rac{1}{2}\theta$
11	aB/aB	Ω	Ω	1	Ω
12	aB/ab	Ω	0	1/2	1/2
13	ab/AB	$rac{1}{2}(1-\theta)$	$rac{1}{2}\theta$	$rac{1}{2}\theta$	$rac{1}{2}(1-\theta)$
14	ab/Ab	0	1/2	0	1/2
15	ab/ab	Ω	0	1/2	1/2
16	ab/ab	0	0	0	1

16 fully phased genotypes can produce are presented in Table [2](#page-3-0) also. For example, entry 2 of Table [2](#page-3-0) shows the genotype of *AB/Ab* and the probabilities of producing the four possible gametes from this genotype are 0.5 for *AB*, 0.5 for *Ab*, 0 for *aB* and 0 for *ab*. Another example is entry 7, which is genotype Ab/ab . This genotype can produce all four gametes with the following probabilities: $\frac{1}{2}\theta$ for *AB*, $\frac{1}{2}(1-\theta)$ for *Ab*, $\frac{1}{2}(1-\theta)$ for *aB* and $\frac{1}{2}\theta$ for *ab*. The 16 \times 4 gametic probability table (the *H* matrix) is the key to form the recurrent equations for genotypes across generations in all mating systems discussed in this study. This *H* matrix can be generated automatically via a computer program.

Recurrent equations of genotype frequencies for self‑fertilization

Starting from the F_2 population with recombination fraction θ , after more than eight generations of continuous self-fertilization, the recombination fraction will reach its equilibrium value [\[6\]](#page-17-7),

The recombination fraction at generation $t < 8$ can be obtained via recurrent equations of genotypes. We will derive the recurrent equations using matrix algebra. Matrix *H* is all what we need to build the 16×16 transition matrix *P*, from which the recurrent equations for computing the frequencies of the 16 genotypes are formed. We denote the genotype frequencies at generation t by a 16×1 vector G_t . The genotypic frequencies at generation $t + 1$ are computed from the frequencies at generation *t*,

$$
G_{t+1} = PG_t \tag{5}
$$

for $t \geq 1$, where P is the transition matrix. The sequences of *G* across generations forms a Markov chain with transition matrix *P*. The above recurrent equations can be further manipulated into

$$
G_{t+1} = PG_t = P^2 G_{t-1} = P^3 G_{t-2} = \dots = P^t G_1 \quad (6)
$$

The genotype frequencies are functions of the genotype frequencies of the initial population (the F_1 individual) with genotype AB/ab and ab/AB , which are the 4th and the 13th genotypes (see Table [2\)](#page-3-0). Therefore, the initial genotype frequency vector has all elements being zero except $G_1[4] = G_1[13] = 1/2.$

We now build the 16×16 transition matrix *P* one column at a time via matrix algebra and through computer programming. Let P_k be the k th column of matrix P for $k = 1, \dots, 16$. Let h_k be the *k*th row of matrix *H* for $k = 1, \dots, 16$ (Table [2\)](#page-3-0). The *k*th column of matrix *P* is

$$
P_k = \text{vec}(h_k^T h_k) \tag{7}
$$

where $\text{vec}(X)$ is a vectorization operator for matrix *X*. For example, if

$$
X = \begin{bmatrix} 1 & 2 & 3 \\ 4 & 5 & 6 \end{bmatrix}
$$

then

$$
vec(X) = \begin{bmatrix} 1 & 4 & 2 & 5 & 3 & 6 \end{bmatrix}^T
$$

which is a column vector. Let us use the following three genotypes as examples to demonstrate the three columns of matrix *P*. For the frst genotype (entry 1 of Table [2](#page-3-0)), we generate the following matrix,

$$
h_1^T h_1 = \begin{bmatrix} 1 \\ 0 \\ 0 \\ 0 \end{bmatrix} \begin{bmatrix} 1 & 0 & 0 & 0 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{bmatrix}
$$

Similarly, we can generate the third genotype (entry 3 of Table [2](#page-3-0)) as

$$
h_3^T h_3 = \begin{bmatrix} \frac{1}{2} \\ 0 \\ 0 \\ \frac{1}{2} \\ 0 \end{bmatrix} \begin{bmatrix} \frac{1}{2} & 0 & \frac{1}{2} & 0 \end{bmatrix} = \begin{bmatrix} \frac{1}{4} & 0 & \frac{1}{4} & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ \frac{1}{4} & 0 & \frac{1}{4} & 0 \\ 0 & 0 & 0 & 0 \end{bmatrix}
$$

and the seventh genotype (entry 7 of Table [2](#page-3-0)) as

$$
h_7^T h_7 = \begin{bmatrix} \frac{1}{2}\theta \\ \frac{1}{2}(1-\theta) \\ \frac{1}{2}(1-\theta) \\ \frac{1}{2}\theta \end{bmatrix} \begin{bmatrix} \frac{1}{2}\theta & \frac{1}{2}(1-\theta) & \frac{1}{2}(1-\theta) & \frac{1}{2}\theta \end{bmatrix} = \begin{bmatrix} \frac{1}{4}\theta^2 & \frac{1}{4}\theta(1-\theta) & \frac{1}{4}\theta(1-\theta) & \frac{1}{4}\theta^2 \\ \frac{1}{4}\theta(1-\theta) & \frac{1}{4}(1-\theta)^2 & \frac{1}{4}(1-\theta)^2 & \frac{1}{4}\theta(1-\theta) \\ \frac{1}{4}\theta(1-\theta) & \frac{1}{4}(1-\theta)^2 & \frac{1}{4}(1-\theta)^2 & \frac{1}{4}\theta(1-\theta) \\ \frac{1}{4}\theta^2 & \frac{1}{4}\theta(1-\theta) & \frac{1}{4}\theta(1-\theta) & \frac{1}{4}\theta^2 \end{bmatrix}
$$

All the $h_k^T h_k$ matrices for $k = 1, \dots, 16$ will be generated this way. From the $h_k^T h_k$ matrix, we build the *k*th column of the 16×16 transition matrix P. For the three genotypes demonstrated above, we obtain the following three column vectors,

Robbins [[16\]](#page-17-3) pooled the 16 fully phased genotypes into 10 genotypes and developed a 10×10 transition matrix. His approach was presented in Supplementary Note S1 for completeness of the study. Haldane and Waddington [[6\]](#page-17-7) further pooled the genotypes into fve classes and

P·¹ = vec(h^T ¹ h1) = 1 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 , P·³ = vec(h^T ³ h3) = 1 4 0 1 4 0 0 0 0 0 1 4 0 1 4 0 0 0 0 0 , P·⁷ = vec(h^T ⁷ h7) = 1 4 θ 2 1 ⁴ θ (1 − θ) 1 ⁴ θ (1 − θ) 1 4 θ 2 1 ⁴ θ (1 − θ) 1 ⁴ (¹ [−] θ)² 1 ⁴ (¹ [−] θ)² 1 ⁴ θ (1 − θ) 1 ⁴ θ (1 − θ) 1 ⁴ (¹ [−] θ)² 1 ⁴ (¹ [−] θ)² 1 ⁴ θ (1 − θ) 1 4 θ 2 1 ⁴ θ (1 − θ) 1 ⁴ θ (1 − θ) 1 4 θ 2

The 16 column vectors form the transition matrix *P*,

$$
P = [P_1 \ P_2 \ P_3 \ \cdots \ P_{16}]
$$

which is given in Supplementary Table S1. Once we fnd the genotypic frequencies using Eq. ([5\)](#page-3-1), we can fnd the recombination fraction at generation *t* by

$$
\theta_{t+1} = W^T G_{t+1} = W^T P^t G_1 \tag{8}
$$

where *W* is a 16×1 vector of weights that are given by the last column of Table [3](#page-5-0). As the number of generations increases, the recombination fraction reaches its limit,

$$
\lim_{t \to \infty} \theta_{t+1} = \lim_{t \to \infty} W^T P^t G_1 = \rho_{\text{self}} = \frac{2\theta}{1 + 2\theta} \tag{9}
$$

For example, when $\theta = 0.1$ in the F₂ population, the fnal recombination fraction in the limit is

$$
\rho_{\text{self}} = \frac{2\theta}{1+2\theta} = \frac{2 \times 0.1}{1+2 \times 0.1} = \frac{1}{6} = 0.166667 \quad (10)
$$

developed a 5×5 transition matrix. Their result is presented in Supplementary Note S2.

Recurrent equations for brother‑sister mating

Recombinant inbred lines generated from brother sister mating is much more complicated than from self-fertilization. Haldane and Waddington [\[6](#page-17-7)] provided the recurrent equations of genotypes and derived the asymptotic solution for the recombination fraction when $t = \infty$, which is

$$
\rho_{\rm sib} = \frac{4\theta}{1 + 6\theta} \tag{11}
$$

Each sibling can take one of the 16 possible fully phased genotypes. Therefore, a sib pair can have a total of $16 \times 16 = 256$ genotype combinations. If we ignore the phase information, there are 10 possible genotypes per individual $[16]$ $[16]$, a sib pair can take one of the $10 \times 10 = 100$ possible genotypes. Haldane and Waddington [[6\]](#page-17-7) pooled the 100 genotypes into 22 composite

Table 3 Recombinant gamete probabilities from all 16 fully phased genotypes and the sum of the two columns used as weights

Entry	Genotype	Pr(Ab)	Pr(aB)	$W = Pr(Ab) + Pr(aB)$
1	AB/AB	0	Ω	0
2	AB/Ab	1/2	0	1/2
3	AB/AB	0	1/2	1/2
4	AB/ab	$rac{1}{2}\theta$	$rac{1}{2}\theta$	θ
5	Ab/AB	1/2	0	1/2
6	Ab/Ab	1	$\mathbf 0$	1
7	Ab/ab		$\frac{1}{2}(1-\theta) = \frac{1}{2}(1-\theta)$	$1-\theta$
8	Ab/ab	1/2	0	1/2
9	aB/AB	0	1/2	1/2
10	aB/Ab		$\frac{1}{2}(1-\theta) = \frac{1}{2}(1-\theta)$	$1-\theta$
11	aB/aB	Ω	1	1
12	aB/ab	0	1/2	1/2
13	ab/AB	$rac{1}{2}\theta$	$rac{1}{2}\theta$	θ
14	ab/Ab	1/2	0	1/2
15	ab/aB	0	1/2	1/2
16	ab/ab	0	0	Ω

genotypes and developed recurrent equations for the 22 composite genotypes at generation $t + 1$ from the frequencies at generation *t*.

We now take advantage of the computer program to generate the 256×256 transition probability matrix and calculate the frequencies of the 256 pairs of genotypes of the sib-pairs. To build the recurrent equations, we frst need to arrange the 16 possible genotypes of the frst sib in the same way as shown in Table [2](#page-3-0). We now nest the second sib's 16 genotypes within each of the frst sib. After defining the order of the sib-pair genotypes in G_t (a 256 \times 1 vector), we are ready to define the transition for $k = 1, \dots, 256$ and $i, j = 1, \dots, 16$. For example, when $i = 4$, $j = 10$, the subscript of P_k is

$$
k = (i - 1) \times 16 + j = (4 - 1) \times 16 + 10 = 48 + 10 = 58
$$

The *k*th column of matrix *P* is

$$
P_{.k} = \text{vec}(h_j^T h_i) \otimes \text{vec}(h_j^T h_i)
$$
\n(13)

We now demonstrate the formation of a few columns of the transition matrix. First, let us demonstrate the second sib-pair, *AB/AB* vs. *AB/Ab*. The gamete probabilities of the sib pair are h_1 and h_2 , respectively. Let us define

$$
h_2^T h_1 = \begin{bmatrix} \frac{1}{2} \\ \frac{1}{2} \\ 0 \\ 0 \end{bmatrix} \begin{bmatrix} 1 & 0 & 0 & 0 \end{bmatrix} = \begin{bmatrix} \frac{1}{2} & 0 & 0 & 0 \\ \frac{1}{2} & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{bmatrix}
$$

Therefore, the vectorization of $h_j^T h_i$ is

$$
\text{vec}(h_2^T h_1) = \left[\frac{1}{2} \frac{1}{2} 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 \right]^T
$$

The column of the transition matrix corresponding to $i = 1$ and $j = 2$ is

$$
k = (i - 1) \times 16 + j = (1 - 1) \times 16 + 2 = 2
$$

Therefore, the 2nd column of matrix P is

$$
P_{.2} = \text{vec}(h_2^T h_1) \otimes \text{vec}(h_2^T h_1) \tag{14}
$$

Let us now illustrate another sib pair, *AB/ab* vs. *Ab/ aB*, where the first sib corresponds to entry number $i = 4$ and the second sib corresponds to entry number $j = 10$. The sib-pair corresponds to column number

$$
k = (i - 1) \times 16 + j = (4 - 1) \times 16 + 10 = 58
$$

of the transition matrix. We frst defne

$$
h_{10}^T h_4 = \begin{bmatrix} \frac{1}{2}\theta \\ \frac{1}{2}(1-\theta) \\ \frac{1}{2}(1-\theta) \\ \frac{1}{2}\theta \end{bmatrix} \begin{bmatrix} \frac{1}{2}(1-\theta) & \frac{1}{2}\theta & \frac{1}{2}(1-\theta) \end{bmatrix} = \begin{bmatrix} \frac{1}{4}\theta(1-\theta) & \frac{1}{4}\theta^2 & \frac{1}{4}\theta^2 & \frac{1}{4}\theta(1-\theta) \\ \frac{1}{4}(1-\theta)^2 & \frac{1}{4}\theta(1-\theta) & \frac{1}{4}(1-\theta)^2 \\ \frac{1}{4}(1-\theta)^2 & \frac{1}{4}\theta(1-\theta) & \frac{1}{4}\theta(1-\theta) & \frac{1}{4}(1-\theta)^2 \\ \frac{1}{4}\theta(1-\theta) & \frac{1}{4}\theta^2 & \frac{1}{4}\theta^2 & \frac{1}{4}\theta(1-\theta) \end{bmatrix}
$$

probability table P (a 256 \times 256 matrix). Recall that the last four columns of Table [2](#page-3-0) form a $16 \times 4H$ matrix. This matrix is also the basic element to develop the transition probability matrix. First, we need to defne the sib pair in the position of vector G_t . If the first sib is entry *i* and the second sib is entry *j*, for $i, j = 1, \dots, 16$, the corresponding sib pair position in vector G_t is defined as

$$
k = (i - 1) \times 16 + j \tag{12}
$$

We then form the 58th column vector of matrix *P*,

$$
P_{.58} = \text{vec}(h_{10}^T h_4) \otimes \text{vec}(h_{10}^T h_4)
$$

We start from the frst column of matrix *P* to the last column of *P* to complete all 256 columns of the *P* matrix, i.e.,

$$
P = [P_{\cdot 1} P_{\cdot 2} \cdots P_{\cdot 256}]
$$

The frequencies of the 256 sib pair genotypes at generation *t* are then used to calculate the frequencies of the sib pair combination for generation $t + 1$, as shown below,

$$
G_{t+1} = PG_t = P^t G_1 \tag{15}
$$

How do we determine the initial sib-pair frequencies? Assume that the initial population is the F_1 hybrid, which represents entries of $i = 4$ (*AB*/*ab*) and $j = 13$ (ab/AB) as shown in Table [2.](#page-3-0) Therefore, the corresponding sib pairs among all $16 \times 16 = 256$ sib-pairs with both sibs being F_1 hybrids are identified as the following four entries,

$$
k_1 = (i - 1) \times 16 + i = (4 - 1) \times 16 + 4 = 52
$$

\n
$$
k_2 = (i - 1) \times 16 + j = (4 - 1) \times 16 + 13 = 61
$$

\n
$$
k_3 = (j - 1) \times 16 + i = (13 - 1) \times 16 + 4 = 196
$$

\n
$$
k_4 = (j - 1) \times 16 + j = (13 - 1) \times 16 + 13 = 205
$$

Therefore, the initial sib-pairs frequencies are

$$
G_1[52] = G_1[61] = G_1[196] = G_1[205] = 1/4
$$

and

$$
G_1[k] = 0, \forall k \not\ni k_1, k_2, k_3, k_4
$$

Recall that the last column of Table [3](#page-5-0) forms a 16×1 weight vector denoted by W . We now build two vectors from vector *W*. The first one is

$$
V_1=W\otimes J_{16\times 1}
$$

and the second one is

$$
V_2 = J_{16 \times 1} \otimes W
$$

where $J_{16\times1}$ is a 16 \times 1 unity vector (all 16 elements are ones) and $X \otimes Y$ is the Kronecker product between matrices X and Y . The final weight vector is the average of the two, i.e.,

$$
V = \frac{1}{2}(V_1 + V_2)
$$
\n(16)

which forms a new 256×1 vector of weights to calculate the recombination fraction at generation $t + 1$.

$$
\theta_{t+1} = V^T G_{t+1} = V^T P^t G_1 \tag{17}
$$

As the number of generations of sib-mating increases, the recombination fraction reaches its limit,

$$
\lim_{t \to \infty} \theta_{t+1} = \lim_{t \to \infty} V^T P^t G_1 = \rho_{\text{sib}} = \frac{4\theta}{1 + 6\theta} \tag{18}
$$

For example, if $\theta = 0.1$, the final recombination fraction in the limit is

$$
\rho_{\rm sib} = \frac{4\theta}{1+6\theta} = \frac{4 \times 0.1}{1+6 \times 0.1} = \frac{1}{4} = 0.25
$$
 (19)

Recurrent equations of gametic frequencies in random mating

Random mating occurs starting from the F_1 hybrid to generate the F_2 and subsequent generations. Such a population is called the advanced intercross lines (AIL) by Darvasi and Soller $[4]$ $[4]$. The advantage of AILs is that linkage between tightly linked loci can be broken thereby increasing recombination. This results in a so-called expanded genetic map where estimated distances appear larger than those of the initial intercross. Such a particular design is useful to order genes or markers in strong linkage at the same locus. For instance, AILs were used for fine mapping in plant genetics $[1]$ $[1]$ and animal genetics [[11\]](#page-17-10). When $t \to \infty$, the recombination fraction reaches the limit, that is $1/2$. Therefore, QTL mapping can be done when t is not too large. There are several different ways to derive the recurrent equations for the recombination fraction. Robbins' [[16\]](#page-17-3) derivation is general so that the initial genotype can be any of the 16 possible genotypes while the derivation of Darvasi and Soller [[4](#page-17-6)] is simple but only applies to the initial genotype of *AB/ab*. The recombination fraction at t can be expressed as a function of the recombination fraction at the F_2 generation.

$$
\theta_t = \frac{1}{2} \Big[1 - (1 - 2\theta)(1 - \theta)^{t - 2} \Big] \tag{20}
$$

for $t \ge 2$. One can verify that when $t = 2$, $\theta_2 = \theta$, which is indeed the recombination fraction of the F_2 population. Denote the four gametic frequencies by a row vector,

$$
G = [p \ q \ s \ t] \tag{21}
$$

where $p = Pr(AB)$, $q = Pr(Ab)$, $s = Pr(aB)$, and $t = Pr(ab)$. Let

$$
G_t = \left\lfloor \begin{array}{c} p_t \ q_t \ s_t \ t_t \end{array} \right\rfloor \tag{22}
$$

be a 1×4 frequency vector of the four gametes at generation *t*. The recurrent equations of the gametic frequencies are

$$
G_{t+1} = (G_t \otimes G_t)H
$$
\n(23)

where *H* is the 16 \times 4 matrix given in Table [2](#page-3-0). For the F₂ population, the four gametic frequencies are

$$
G_1 = \left[\frac{1}{2}(1-\theta) \frac{1}{2}\theta \frac{1}{2}\theta \frac{1}{2}(1-\theta)\right]
$$
 (24)

In contrast to the previous mating systems, the gametic frequencies at generation $t + 1$ for random mating are

Fig. 1 Recombination fraction profles after 10 generations of self-fertilization

Table 4 Asymptotic recombination fractions from diferent initial recombination fractions after repeated self-fertilization for 10 generations

not linear functions of the gametic frequencies at generation t . The recombination fraction for generation t is

$$
\theta_t = G_t[2] + G_t[3] = q_t + s_t \tag{25}
$$

If the initial gametic frequency vector is the one given in Eq. [\(24](#page-6-0)), the limit of θ_t is 0.5 as $t \to \infty$. Using the result of Robbins $[16]$ $[16]$, we can prove Eq. (20) (20) , which is presented in Supplementary Note S3.

Results

Self fertilization

We frst demonstrate the recombination fraction trajectory across generations when self-fertilization starts from the F_1 hybrid, i.e., the initial genotype frequencies for AB/ab and ab/AB are 1/2 and 1/2, and 0 for all other genotypes. The initial recombination fraction was set at the following levels: 0.05, 0.1, 0.15, 0.2, 0.25 and 0.3. We

evaluate the trajectory of recombination fractions for 10 generations, as shown in Fig. [1](#page-7-0). After 10 generations of self-fertilization, they all reach their asymptotic values, which are presented in Table [4](#page-7-1). For example, when the initial recombination fraction is $\theta = 0.20$, the asymptotic recombination fraction is

$$
\rho_{Self} = \frac{2\theta}{1+2\theta} = \frac{2 \times 0.20}{1+2 \times 0.20} = 0.285714 \tag{26}
$$

Brother‑sister mating

Among all $16 \times 16 = 256$ sib-pairs with both sibs being F_1 hybrids are identified as the following four entries,

$$
k_1 = (i - 1) \times 16 + i = (4 - 1) \times 16 + 4 = 52
$$

\n
$$
k_2 = (i - 1) \times 16 + j = (4 - 1) \times 16 + 13 = 61
$$

\n
$$
k_3 = (j - 1) \times 16 + i = (13 - 1) \times 16 + 4 = 196
$$

\n
$$
k_4 = (j - 1) \times 16 + j = (13 - 1) \times 16 + 13 = 205
$$

Therefore, the initial frequencies are

$$
G_0[52] = G_0[61] = G_0[196] = G_0[205] = 1/4
$$

and

$$
G_0[k] = 0, \forall k \notin k_1, k_2, k_3, k_4
$$

Again, we demonstrate the recombination fraction profles across generations for brother-sister mating starting from the F_1 population. The four sib pairs corresponding to the double heterozygote are (AB/ab-AB/ ab), (AB/ab-ab/AB), (ab/AB-AB/ab) and (ab/AB-ab/ AB). The initial recombination fraction was set at the

Fig. 2 Recombination fraction profles after 20 generations of brother-sister mating

following levels: 0.05, 0.1, 0.15, 0.2, 0.25 and 0.3. We evaluated the recombination fractions change for 20 generations. Figure [2](#page-8-0) shows the recombination fraction trajectories. After 20 generations of brother-sister mating, they all reach their equilibrium values, which are presented in Table [5](#page-8-1). For example, when the initial recombination fraction is $\theta = 0.20$, the asymptotic recombination fraction is

Random mating

Starting from the F_1 hybrid, the population went to 50 generations of random mating. The recombination fraction profles are demonstrated in Fig. [3](#page-9-0) from various initial values of the recombination fraction.

After 50 generations of random mating, all populations have reached their equilibrium value ($\rho_{Random} = 0.5$) except that the population starting with $\theta = 0.05$ has not reached the equilibrium, but with a recombination fraction of 0.4653748 at $t = 50$, which is calculated via

$$
\theta_{50} = \frac{1}{2} \Big[1 - (1 - 2\theta)(1 - \theta)^{50 - 2} \Big] = \frac{1}{2} \Big[1 - (1 - 2 \times 0.05)(1 - 0.05)^{48} \Big] = 0.4616 \tag{28}
$$

$$
\rho_{Sib} = \frac{4\theta}{1+6\theta} = \frac{4 \times 0.20}{1+6 \times 0.20} = 0.3636364 \qquad (27)
$$

Table 5 Asymptotic recombination fractions from diferent initial recombination fractions after repeated brother-sister mating for 20 generation

Three R functions were provided to calculate the recombination fractions for PRERILs developed via selffertilization, brother-sister mating and random mating (Supplementary Code S1).

Comparison of the three mating systems

Starting with the same recombination fraction of 0.10, all three mating systems (self-fertilization, brother-sister mating, and random mating) underwent 20 consecutive generations of reproduction. The trajectories of the recombination fraction are compared among the three mating systems (Fig. [4](#page-9-1)).

At generation 2 and 3, all three mating systems have the same recombination fraction. Self-fertilization starts splitting from the other two systems after generation 3 while brother-sister mating deviates from random mating after generation 4. Four generations of self-fertilization

Fig. 3 Recombination fraction profles after 50 generations of random mating

Fig. 4 Comparison of recombination fraction profles among three mating systems

are technically sufficient to make the recombination fraction reach the asymptotic value. Ten generations of brother-sister mating are sufficient to make the recombination fraction reach its equilibrium value.

Validation of the recurrent equations via Monte Carlo simulations

The recurrent equations derived here must be correct because the asymptotic results $(t > 10$ for self-fertilization and $t > 20$ for brother-sister mating) match the final results provided by Haldane and Waddington [\[6](#page-17-7)] and

Robbins [[14\]](#page-17-4). To doubly ensure the correctness of the derivation, we decided to further validate the recurrent equations via Monte Carlo simulations. The assumptions required to derive the recurrent equations are (1) there is no interference in crossovers between two genomic segments [[5\]](#page-17-11); (2) there is no segregation distortion of the markers under investigation. As a result, the best way to validate the recurrent equations is to simulate RILs based on these assumptions. It is hard to use actual data from RILs for validation because these two assumptions may not be satisfed in reality. Another justifcation for using simulations to validate the derivation is that the derivation is based on expectations of the genotype groups and the expectations only apply to large samples, in fact, infnitely large samples. In reality, the sample sizes of real populations are always finite. The theoretical derivation cannot be validated based on one or a few fnite samples.

The recurrent equation of the recombination fraction for random matting was originally derived by Jennings [[9\]](#page-17-0) and later by Darvasi and Soller [[4\]](#page-17-6). Our derivation is merely an alternative approach to obtain the same result. Therefore, no validation is needed for random mating. To validate the recurrent equations for self-fertilization, we started with a single F1 individual with genotype AB/ ab. We replaced this phased genotype by 11/00, where 11 is the paternal haplotype and 00 is the maternal haplotype in the F1 founder. The distance between the two loci set at

and the second locus of the paternal allele has passed to the gamete, i.e., 0. If x is an even number, recombination would not happen and thus the second locus of the gamete would remain 1 from the paternal allele. The same algorithm also applied to the maternal haplotype of the gamete. The two gametes merged together to make the genotype of the individual. This process continued until $t = 10$ generations. We generated 500 individuals from the single seed descent process to make up one RIL population. The recombination fraction at generation t was the proportion of the recombinant haplotypes, 10 or 01. Note that the two recombinant haplotypes are referred to the $F₂$ generation. In later generations, 11 and 00 may be the recombinants. Eventually, we generated 20 populations, each consisting of 500 individuals. Figure [5](#page-10-0) (panel A) shows the recombination fractions of each population against the generation index up to 10. Variation among

$$
\mu = -\frac{1}{2}\ln(1 - 2\theta) = -\frac{1}{2}\ln(1 - 2 \times 0.1) = 0.1115718 \text{ Morgan}
$$
\n(29)

where the recombination fraction was set at $\theta = 0.1$. The simulation started at F2 from the F1 gametes. There were two random numbers involved in generating each gamete. The first random number was Bernoulli δ_1 ~ Bernoulli(0.5), which determines the first allele of the gamete. If $\delta_1 = 1$ then the gamete took 1 from the paternal allele; otherwise, the gamete took 0 from the maternal allele. Let us assume that $\delta_1 = 1$ so that the paternal allele has passed to the gamete for the frst locus. We then generated a Poisson random from mean $\mu = 0.1115718$, say $x = 0, 1, \dots, \infty$, i.e., $x \sim \text{Poisson}(\mu)$. If x is an odd number, then recombination has occurred

the 20 replicates is very obvious. The average of the 20 replicates is shown in the scatter plot, which partially overlaps with theoretical curve in blue. Figure [5](#page-10-0) (panel B) shows the same plots but the sample size of each population has been increased to 10,000. The average of the 20 populations (scatter plots) completely overlaps with the blue theoretical value. Even though the sample size was as large as 10,000, there were still variation among the replicates.

We also simulated brother-sister mating for 20 generations with sample size of 500 mating pairs and 10,000 mating pairs, respectively. Both schemes were replicated

Fig. 5 Simulation results of self-fertilization for 10 generations. Panel **A** shows 20 replications of sample size 500. Panel **B** shows 20 replications of sample size 10,000. The blue curves are the theoretical values from the recurrent equations while the scatter plots show the averages of the 20 replications

Fig. 6 Simulation results of brother-sister mating for 20 generations. Panel **A** shows 20 replications for sample size of 500 mating pairs. Panel **B** shows 20 replications for sample size of 10,000 mating pairs. The blue curves are the theoretical values from the recurrent equations while the scatter plots show the averages of the 20 replications

20 times. The results are shown in Fig. [6,](#page-11-0) where Fig. [6A](#page-11-0) shows the plots for sample size of 500 mating pairs and Fig. [6B](#page-11-0) shows the results for sample size of 10,000 mating pairs. The variation among the 20 replicates was smaller than the variation in self-fertilization because the sample size was actually doubled in brother-sister mating. The simulation studies have successfully validated the theoretical derivation of the recurrent equations.

Incorporation of the modifed recombination fraction in QTL mapping

Starting with $\theta = 0.05$, if the F₁ hybrid undergoes 3 generations of self-fertilization, the recombination fraction will change from $\theta = 0.05$ in F₂ to $\theta_4 = 0.08923$ in F₄. The heterozygosity will be reduced from 0.5 to

QTL mapping using F_4 as the source population requires a new approach to calculate the conditional probabilities of QTL genotypes given fanking marker genotypes. This is due to (1) The recombination fraction between two loci has been modified from θ in the initial population to θ_4 in the F₄ population; (2) The heterozygosity has been modified from $H_2 = 0.50$ to $H_4 = 0.125$. Below is an example showing the diferences in QTL mapping between F_2 and F_4 .

*QTL mapping in F***2** *populations*

For the F_2 population, let the recombination fraction between the two fanking markers (A and B) be θ_{AB} = 0.05, the recombination fraction between A and Q be $\theta_{AQ} = 0.01$ and thus the recombination fraction between Q and B is (31)

$$
\theta_{QB} = (\theta_{AB} - \theta_{AQ})/(1 - 2\theta_{AQ}) = (0.05 - 0.01)/(1 - 2 \times 0.01) = 0.0408
$$

where the order of the three loci is A-Q-B. Let us denote the genotypes of the three loci by

$$
A = \begin{cases} 1 & \text{for AA} & \text{Pr}(AA) = 0.25 \\ 2 & \text{for A}a & \text{Pr}(Aa) = 0.50 \\ 3 & \text{for aa} & \text{Pr}(aa) = 0.25 \end{cases}; B = \begin{cases} 1 & \text{for BB} & \text{Pr}(BB) = 0.25 \\ 2 & \text{for BB} & \text{Pr}(BB) = 0.50 \\ 3 & \text{for bb} & \text{Pr}(bb) = 0.50 \end{cases}; Q = \begin{cases} 1 & \text{for QQ} & \text{Pr}(QQ) = 0.25 \\ 2 & \text{for QQ} & \text{Pr}(Qq) = 0.50 \\ 3 & \text{for qq} & \text{Pr}(qq) = 0.25 \end{cases}
$$

$$
H_4 = \left(\frac{1}{2}\right)^{4-1} = \frac{1}{8} = 0.125\tag{30}
$$

The conditional probabilities of the QTL genotype given the fanking marker genotypes is defned from the following Bayes theorem,

$$
Pr(Q = k | A = i, B = j) = \frac{Pr(Q = k) Pr(A = i | Q = k) Pr(B = j | Q = k)}{\sum_{k'=1}^{3} Pr(Q = k') Pr(A = i | Q = k') Pr(B = j | Q = k')} \tag{32}
$$

where $Pr(Q = k)$ is the marginal probability of the QTL genotype, $Pr(A = i | Q = k)$ and $Pr(B = j | Q = k)$ are the conditional probabilities of the marker genotypes given the QTL genotype. The conditional probabilities are extracted from the following two transition matrices,

$$
T_{A/Q} = \begin{bmatrix} Pr(A = 1|Q = 1) & Pr(A = 2|Q = 1) & Pr(A = 3|Q = 1) \\ Pr(A = 1|Q = 2) & Pr(A = 2|Q = 2) & Pr(A = 3|Q = 2) \\ Pr(A = 1|Q = 3) & Pr(A = 2|Q = 3) & Pr(A = 3|Q = 3) \end{bmatrix}
$$

$$
= \begin{bmatrix} (1 - \theta_{QA})^2 & 2(1 - \theta_{QA})\theta_{QA} & \theta_{QA}^2 \\ (1 - \theta_{QA})\theta_{QA} & (1 - \theta_{QA})^2 + \theta_{QA}^2 & (1 - \theta_{QA})\theta_{QA} \\ \theta_{QA}^2 & 2(1 - \theta_{QA})\theta_{QA} & (1 - \theta_{QA})^2 \end{bmatrix}
$$

and

$$
T_{B/Q} = \begin{bmatrix} Pr(B = 1|Q = 1) & Pr(B = 2|Q = 1) & Pr(B = 3|Q = 1) \ Pr(B = 1|Q = 2) & Pr(B = 2|Q = 2) & Pr(B = 3|Q = 2) \ Pr(B = 1|Q = 3) & Pr(B = 2|Q = 3) & Pr(B = 3|Q = 3) \end{bmatrix}
$$

$$
= \begin{bmatrix} (1 - \theta_{QB})^2 & 2(1 - \theta_{QB})\theta_{QB} & \theta_{QB}^2 \ (1 - \theta_{QB})\theta_{QB} & (1 - \theta_{QB})^2 + \theta_{QB}^2 & (1 - \theta_{QB})^2 \end{bmatrix}
$$

For example, the conditional probability that the QTL genotype is Qq given the genotype of locus A is AA and the genotype of locus B is Bb is

Therefore, the three conditional probabilities of the QTL genotypes are

$$
Pr(Q = 1|A = 1, B = 2) = \frac{(1 - \theta_{QA})^2 \theta_{QB} (1 - \theta_{QB})}{\theta_{AB} (1 - \theta_{AB})} = 0.807816
$$
\n(37)\n
$$
Pr(Q = 2|A = 1, B = 2) = \frac{(1 - \theta_{QA})\theta_{QA} \left[(1 - \theta_{QB})^2 + \theta_{QB}^2 \right]}{\theta_{AB} (1 - \theta_{AB})}
$$
\n= 0.1921015\n(38)

$$
Pr(Q = 3|A = 1, B = 2) = \frac{\theta_{QA}^2 \theta_{QB} (1 - \theta_{QB})}{\theta_{AB} (1 - \theta_{AB})} = 0.0000824 \quad (39)
$$

Here is another example, the conditional probability that the QTL genotype is Qq given the genotype of locus A is AA and the genotype of locus B is BB is

$$
Pr(Q = 2|A = 1, B = 1)
$$

=
$$
\frac{Pr(Q = 2)Pr(A = 1|Q = 2)Pr(B = 1|Q = 2)}{\sum_{k'=1}^{3} Pr(Q = k')Pr(A = 1|Q = k')Pr(B = 1|Q = k')}
$$

(40)

$$
Pr(Q = 2|A = 1, B = 2) = \frac{Pr(Q = 2)Pr(A = 1|Q = 2)Pr(B = 2|Q = 2)}{\sum_{k'=1}^{3} Pr(Q = k')Pr(A = 1|Q = k')Pr(B = 2|Q = k')} \tag{33}
$$

where

$$
Pr(Q=1)Pr(A=1|Q=1)Pr(B=2|Q=1) = \frac{1}{4}(1-\theta_{QA})^2 \times 2\theta_{QB}(1-\theta_{QB}) = 0.0191856
$$
\n(34)

$$
Pr(Q=2) Pr(A=1|Q=2) Pr(B=2|Q=2) = \frac{1}{2}(1-\theta_{QA})\theta_{QA} \left[(1-\theta_{QB})^2 + \theta_{QB}^2 \right] = 0.0045624
$$
\n(35)

$$
Pr(Q=3) Pr(A=1|Q=3) Pr(B=2|Q=3) = \frac{1}{4} \theta_{QA}^2 \times 2\theta_{QB} (1 - \theta_{QB}) = 1.9575E - 6
$$
\n(36)

The denominator is the sum of the three terms,

$$
\sum_{k'=1}^{3} \Pr(Q = k') \Pr(A = 1 | Q = k') \Pr(B = 2 | Q = k') = \frac{1}{2} \theta_{AB} (1 - \theta_{AB}) = 0.02375
$$

 (26)

where

$$
Pr(Q = 1) Pr(A = 1|Q = 1) Pr(B = 1|Q = 1) = \frac{1}{4}(1 - \theta_{QA})^2(1 - \theta_{QB})^2
$$
\n(41)

$$
Pr(Q = 2) Pr(A = 1|Q = 2) Pr(B = 1|Q = 2) = \frac{1}{2}(1 - \theta_{QA})\theta_{QA}(1 - \theta_{QB})\theta_{QB}
$$
\n(42)

$$
Pr(Q = 3) Pr(A = 1|Q = 3) Pr(B = 1|Q = 3) = \frac{1}{4} \theta_{QB}^2 \theta_{QA}^2 \tag{43}
$$

The denominator is the sum of the three terms,

$$
\frac{1}{4}(1 - \theta_{QA})^2 (1 - \theta_{QB})^2 + \frac{1}{2}(1 - \theta_{QA})\theta_{QA}(1 - \theta_{QB})\theta_{QB} + \frac{1}{4}\theta_{QB}^2 \theta_{QA}^2
$$
\n
$$
= \frac{1}{4}[(1 - \theta_{QA})(1 - \theta_{QB}) + \theta_{QB}\theta_{QA}]^2
$$
\n
$$
= \frac{1}{4}[1 - (\theta_{QB} + \theta_{QA} - 2\theta_{QA}\theta_{QB})]^2
$$
\n
$$
= \frac{1}{4}(1 - \theta_{AB})^2
$$
\n(44)

Therefore, the three conditional probabilities of genotypes given the genotypes of markers A and B are

$$
= 0.07296
$$
\nwhere the order of the three loci is A-Q-B. Let us denote the genotypes of the three loci by

\nthe QTL

 $\theta_{OB}^{(4)} = (\theta_{AB}^{(4)} - \theta_{AO}^{(4)})/(1 - 2\theta_{AO}^{(4)})$

between Q and B in F_4 is

$$
A = \begin{cases} 1 & \text{for AA} & \text{Pr}(AA) = 0.4375 \\ 2 & \text{for A}a & \text{Pr}(Aa) = 0.125 \\ 3 & \text{for aa} & \text{Pr}(aa) = 0.4375 \end{cases}; B = \begin{cases} 1 & \text{for BB} \\ 2 & \text{for BB} \\ 3 & \text{for bb} \end{cases}
$$

$$
Pr(Q = 1|A = 1, B = 1) = \frac{\frac{1}{4}(1 - \theta_{QA})^2(1 - \theta_{QB})^2}{\frac{1}{4}(1 - \theta_{AB})^2}
$$

$$
= \frac{(1 - \theta_{QA})^2(1 - \theta_{QB})^2}{(1 - \theta_{AB})^2}
$$

$$
= 0.9991749
$$
(45)

$$
Pr(Q = 2|A = 1, B = 1) = \frac{2(1 - \theta_{QA})\theta_{QA}(1 - \theta_{QB})\theta_{QB}}{(1 - \theta_{AB})^2}
$$

$$
= \frac{2 \times (1 - 0.01) \times 0.01 \times (1 - 0.0408) \times 0.0408}{(1 - 0.05)^2}
$$

$$
= 0.0008585929
$$
(46)

and

$$
Pr(Q = 3|A = 1, B = 1) = \frac{\frac{1}{4}\theta_{QA}^2 \theta_{QB}^2}{\frac{1}{4}(1 - \theta_{AB})^2}
$$

$$
= \frac{\theta_{QA}^2 \theta_{QB}^2}{(1 - \theta_{AB})^2}
$$

$$
= 1.844476E - 07
$$
(47)

*QTL mapping in F***4** *populations*

For the F_4 population, the recombination fraction between the two fanking markers (A and B) has changed from $Pr(BB) = 0.4375$ $Pr(Bb) = 0.125$ $Pr(bb) = 0.4375$ $; Q =$ $\sqrt{ }$ \overline{I} \mathcal{L} 1 2 3 for for for QQ Qq qq $Pr(QQ) = 0.4375$ $Pr(Qq) = 0.125$ $Pr(qq) = 0.4375$

 $\theta_{AB} = 0.05$ to $\theta_{AB}^{(4)} = 0.08923$, the recombination fraction between A and Q has changed from $\theta_{AQ} = 0.01$ to $\theta_{AQ}^{(4)} = 0.01905$ and thus the recombination fraction

 $= (0.08923 - 0.01905)/(1 - 2 \times 0.01905)$

Note that the superscript of θ is now the generation index because the subscript has been taken by the two loci. The conditional probabilities of the QTL genotype given the fanking marker genotypes is defned from the following Bayes theorem,

$$
\Pr(Q = k | A = i, B = j) = \frac{\Pr(Q = k) \Pr(A = i | Q = k) \Pr(B = j | Q = k)}{\sum_{k'=1}^{3} \Pr(Q = k') \Pr(A = i | Q = k') \Pr(B = j | Q = k')} \tag{49}
$$

where $Pr(Q = k)$ is the marginal probability of the QTL genotype, $Pr(A = i | Q = k)$ and $Pr(B = j | Q = k)$ are the conditional probabilities of the marker genotypes given the QTL genotype.

These transition probabilities are extracted from the following two transition matrices,

$$
T_{A/Q} = \left[\begin{array}{ccc} (1-\theta_{(A)}^{(4)})^2 & 2(1-\theta_{(A)}^{(4)})\theta_{(A)}^{(4)} & \theta_{(A)}^{(4)2} \\ (1-\theta_{(A)}^{(4)})\theta_{(A)}^{(4)} & (1-\theta_{(A)}^{(4)})^2 + \theta_{(A)}^{(4)2} & (1-\theta_{(A)}^{(4)})\theta_{(A)}^{(4)} \\ \theta_{(A)}^{(4)2} & 2(1-\theta_{(A)}^{(4)})\theta_{(A)}^{(4)} & (1-\theta_{(A)}^{(4)})^2 \end{array} \right]
$$

and

$$
T_{B/Q} = \left[\begin{array}{ccc} (1-\theta_{(B)}^{(4)})^2 & 2(1-\theta_{(2B)}^{(4)})\theta_{(2B)}^{(4)} & \theta_{(2B)}^{(4)2} \\ (1-\theta_{(2B)}^{(4)})\theta_{(2B}^{(4)} & (1-\theta_{(2B)}^{(4)})^2 + \theta_{(2B)}^{(4)2} & (1-\theta_{(2B)}^{(4)})\theta_{(2B}^{(4)} \\ \theta_{(2B)}^{(4)2} & 2(1-\theta_{(2B)}^{(4)})\theta_{(2B)}^{(4)} & (1-\theta_{(2B)}^{(4)})^2 \end{array}\right]
$$

(48)

For example, the conditional probability that the QTL genotype is Qq given the genotype of locus A is AA and the genotype of locus B is Bb is

$$
Pr(Q = 3) Pr(A = 1|Q = 3) Pr(B = 1|Q = 3)
$$

= 0.4375 × $\theta_{QB}^{2} \theta_{QA}^{2} = 8.4515E - 7$

$$
Pr(Q = 2|A = 1, B = 2) = \frac{Pr(Q = 2)Pr(A = 1|Q = 2)Pr(B = 2|Q = 2)}{\sum_{k'=1}^{3} Pr(Q = k')Pr(A = 1|Q = k')Pr(B = 2|Q = k')} \tag{50}
$$

where

$$
Pr(Q = 1) Pr(A = 1|Q = 1) Pr(B = 2|Q = 1)
$$

= 0.4375 × (1 - $\theta_{QA}^{(4)}$)² × 2 $\theta_{QB}^{(4)}$ (1 - $\theta_{QB}^{(4)}$)
= 0.0569487

$$
Pr(Q = 2) Pr(A = 1|Q = 2) Pr(B = 2|Q = 2)
$$

= 0.125 × (1 - $\theta_{QA}^{(4)}$) $\theta_{QA}^{(4)}$ [(1 - $\theta_{QB}^{(4)})^2$ + $\theta_{QB}^{(4)2}$]
= 0.0020199

$$
Pr(Q = 3) Pr(A = 1|Q = 3) Pr(B = 2|Q = 3)
$$

= 0.4375 × $\theta_{QA}^{(4)2}$ × $2\theta_{QB}^{(4)}(1 - \theta_{QB}^{(4)})$
= 0.0000215

The denominator is the sum of the three terms,

$$
\sum_{k'=1}^{3} \Pr(Q = k') \Pr(A = 1 | Q = k') \Pr(B = 2 | Q = k') = 0.0589901
$$

Therefore, the three conditional probabilities of the QTL genotypes given the genotypes of markers A and B are

$$
Pr(Q = 1 | A = 1, B = 2) = \frac{0.0569487}{0.0589901} = 0.9653945
$$

$$
Pr(Q = 2|A = 1, B = 2) = \frac{0.0020199}{0.0589901} = 0.0342414
$$

$$
Pr(Q = 3|A = 1, B = 2) = \frac{0.0000215}{0.0589901} = 0.0003641
$$

The conditional probability that the QTL genotype is Qq given the genotype of locus A is AA and the genotype of locus B is BB is

The denominator is the sum of the three terms,

$$
\sum_{k'=1}^{3} \Pr(Q = k') \Pr(A = 1|Q = k') \Pr(B = 1|Q = k') = 0.3619592
$$

Therefore, the three conditional probabilities of the QTL genotypes given the genotypes of markers A and B are

$$
Pr(Q = 1 | A = 1, B = 1) = \frac{0.3618003}{0.3619592} = 0.9995612
$$

$$
Pr(Q = 2|A = 1, B = 1) = \frac{0.000158}{0.3619592} = 0.0004365
$$

and

$$
Pr(Q = 3 | A = 1, B = 1) = \frac{8.4515E - 7}{0.3619592} = 2.3349E - 6
$$

Table [6](#page-15-0) summarizes the results of the conditional probabilities of the QTL genotypes given the fanking marker genotypes. Because of the heterozygosity reduction due to repeated inbreeding, the conditional probabilities of homozygotes in F_4 are always higher than the homozygosity in F_2 .

Validation from pre‑recombinant inbred lines of rice

Xu et al. [\[18](#page-17-12)] generated 191 F_2 plants from an elite hybrid rice (Shanyou63) derived from the cross of Zhenshan97 and Minghui63. From the 191 $F₂$ plants, they further developed 191 F_3 and 191 F_4 plants by single seed descent. Genotypes of a total of 1696 marker bins were inferred from the DNA sequences for each plant of the three filiations. The data set was used to validate the Markov model. Technically, one pair of markers suffice to validate the model. We used

$$
Pr(Q = 2|A = 1, B = 1) = \frac{Pr(Q = 2)Pr(A = 1|Q = 2)Pr(B = 1|Q = 2)}{\sum_{k'=1}^{3} Pr(Q = k')Pr(A = 1|Q = k')Pr(B = 1|Q = k')}
$$

where

$$
Pr(Q = 1) Pr(A = 1|Q = 1) Pr(B = 1|Q = 1) = 0.4375 \times (1 - \theta_{QA})^2 (1 - \theta_{QB})^2 = 0.3618003
$$

$$
Pr(Q = 2) Pr(A = 1|Q = 2) Pr(B = 1|Q = 2) = 0.125 \times (1 - \theta_{OA})\theta_{OA} (1 - \theta_{OB})\theta_{OB} = 0.000158
$$

Table 6 Comparison of the conditional probabilities of QTL genotypes given flanking marker genotypes between F_2 and F_4 generations of self-fertilization

	Prior		Posterior	
	F,	$F_{\scriptscriptstyle A}$	F,	$F_{\scriptscriptstyle A}$
$Pr(Q = 1 A = 1, B = 2)$	0.25	0.4375	0.807816	0.9653945
$Pr(Q = 2 A = 1, B = 2)$	05	0.125	0.1921015	0.0342414
$Pr(Q = 3 A = 1, B = 2)$	025	0.4375	0.0000824	0.0003641
$Pr(Q = 1 A = 1, B = 1)$	025	0.4375	0.9991409	0.9995612
$Pr(Q = 2 A = 1,B = 1)$	05	0.125	0.0008589	0.0004365
$Pr(Q = 3 A = 1, B = 1)$	025	0.4375	1.846F-7	2.3349F-6

markers of the 12th chromosome for validation. This is the shortest chromosome with 63 markers, all of which follow Mendelian segregations in all three fliations (F_2 , F_3 and F_4). The Mendelian ratio for F_2 is 1 : 2 : 1, for F_3 is 3 : 2 : 3 and for F_4 is 7 : 2 : 7, which were used as the theoretical proportions in the segregation distortion Chi-square tests. The 63 markers form $63 \times 62/2 = 1953$ marker pairs for recombination fraction analyses. There were 19 co-segregating marker pairs in F_2 and thus only 1953 – 19 = 1934 pairs of markers were used in the validation tests. Since the true recombination fractions of the marker pairs in the F_2 generation were not known, we did not have the true recombination fractions to start with for calculating the theoretical recombination fractions for the F_3 and F_4 generations. We treated the observed recombination fractions for the F_2 generation as the "true" values to calculate the theoretical recombination fractions of the F_3 and F_4 generations. To reduce the impact of the unknown initial recombination fractions of $F₂$ on the theoretical values of the recombination fractions in F_3 and F_4 , we took the average recombination fraction of all marker pairs with recombination fractions in the neighborhood of 0.05, 0.10, 0.15, 0.20, 0.25, 0.30, 0.35 and 0.40. The predicted recombination fractions of these marker pairs in the F_3 and F_4 generations were compared with the 95% confdence intervals (95% CIs) of the estimated recombination fractions. The estimated recombination fractions and the 95% CIs were calculated using the method described below.

Instead of directly estimating the recombination fractions between two markers using the expectation– maximization (EM) algorithm, we frst estimated the correlation coefficient between the numerically coded genotypes (0, 1 and 2) of the two markers in the marker pair. Denote the estimated correlation coefficient between markers i and j by r_{ij} with a standard error of

$$
s_r = \sqrt{\frac{1 - r_{ij}^2}{n - 2}}\tag{51}
$$

where $n = 191$ is the sample size. The corresponding recombination fraction between the two markers is

$$
\theta_{ij} = \frac{1}{2}(1 - r_{ij})\tag{52}
$$

with a standard error of θ_{ij} is

$$
s_{\theta} = \sqrt{\text{var}(\theta_{ij})} = \sqrt{\frac{1}{4} \text{var}(r_{ij})} = \frac{1}{2} s_r \tag{53}
$$

The asymptotic 95% confidence interval is

$$
\theta_{ij} - 1.96s_{\theta} < \theta_{ij} < \theta_{ij} + 1.96s_{\theta} \tag{54}
$$

Figure [7](#page-16-0) compares the theoretical recombination fractions (solid lines) calculated from the recurrent equations with the 95% confdence bands (light blue areas) of the estimated recombination fractions for F_2 , F_3 and F_4 . The 95% confdence bands cover the theoretical recombinant fractions in all situations except F_4 (the upper right panel) where the theoretical value barely touches the upper bound. The conclusion is that the theoretical recombination fractions calculated from the Markov model are valid.

Discussion

The recurrent equations of genotype frequency array are Markov chains, which consist of two components: the probabilities of multiple states and the transition probabilities. Historically, the smaller the number of states, the easier the calculation. This was the very reason why Robbins [[16\]](#page-17-3) pooled the 16 fully phased genotypes of two loci into 10 unphased genotypes. Haldane and Waddington [\[6](#page-17-7)] further reduced the number of genotypes from 10 to 5. The reduction of the number of genotypes was very important in reducing the computational burdens in the pre-computer age. People can manually derive the transition probability matrix because of the lower dimension of the matrix. In the computer era, everything can be generated with computer code. The reduction of the number of genotypes is no longer important. We are dealing with a problem that the parameter (recombination fraction) is derived with recurrent equations, not estimated from observed data. Therefore, combining high dimensional fully phased genotypes into low dimensional unphased genotypes has lost its advantage. In fact, utilization of the fully phased genotypes with computer code can avoid human errors in manually writing the transition matrix.

Fig. 7 Predicted recombination fractions from the Markov model (solid lines) and the 95% confidence bands (light blue areas) from F_{2} , F_3 and F_4 of a rice population

An essential component of genetic mapping with RILs is to reconstruct the parental origins (the haplotypes) of DNA on the RIL chromosomes. In addition, QTL mapping using RILs as the genetic resources is a common practice in plants and small laboratory animals. With self-fertilization, as few as 8 generations are required [[13\]](#page-17-13). How do we justify QTL mapping with PRERILs vs. RILs? Whether saving just a few of years using PRERILs for QTL mapping compared to using RILs is worth the effort considering the complexity of the mapping procedure. We argue that optimal utilization of the available genetic resources is always a factor to consider. If phenotypes and genotypes of PRERILs are available, why do we want to waste that information? QTL mapping with PRE-RILs may be important for laboratory animals because development of RILs requires about 20 generations of brother-sister mating. If we can perform QTL mapping with PRERILs half-way before RILs are fully developed, the time saved may be significant. The advanced intercross lines (AILs) increase the proportion of recombination between any two loci and thus provide precision to mapping closely linked QTL. The genetic basis of genome-wide association studies (GWAS) comes from the increased recombination fractions between loci.

Another justifcation for the study of recombination fraction in PRERILs is purely for scientifc reason. We knew the recombination fraction both in the beginning (F_2) and in the end (RILs) but did not know the trajectory how it reaches the equilibrium. This study for the first time flls the gap left for over 100 years.

There are many forms of repeated inbreeding. Jennings [[7–](#page-17-14)[9\]](#page-17-0) investigated at least a dozen forms of them, including random mating, parent–offspring mating, assortative mating, self-fertilization, brother-sister mating, and selection with relation to one of the two loci. Robbins [[14–](#page-17-4)[16\]](#page-17-3) reinvestigated majority of Jennings mating systems plus selection of dominants with respect to one of the two linked characters. Haldane and Waddington [\[6](#page-17-7)] investigate self-fertilization and parent–ofspring mating with great details. Among all the mating systems, self-fertilization, and brother-sister mattings are the main forms of inbreeding to generate recombinant inbred lines.

In modern genetics, more advanced breeding systems have been developed for plants and laboratory animals, such as the Multi-parent Advanced Generation Inter-Cross (MAGIC) population in *Arabidopsis thaliana* [\[10](#page-17-15)] and the Collaboratory Crosses (CC) in mice $[3]$ $[3]$. The RILs of mice derived from an 8-way crosses of mice [\[2](#page-17-8)] were extension of the two -way cross of brother-sister mattings. Recurrent equations of genotype array and the recombination fraction between two loci in these complex inbreeding systems are difficult to derive. The number of genotype array can be huge, and the transition matrix may be in the order of thousand or ten thousand. Manual derivation is certainly not an option. If there is an interest, computer programs may be developed in the future to deal with the complex mating systems.

Conclusions

We developed recurrent equations for calculating genotype frequencies for pre-recombinant inbred lines (PRERILs). These equations allow us to compute the recombination fractions between two loci before the lines reach the equilibrium state. An R function is provided for users to calculate the recombination fractions in PRERILs.

Abbreviations

Supplementary Information

The online version contains supplementary material available at [https://doi.](https://doi.org/10.1186/s12864-024-10699-z) [org/10.1186/s12864-024-10699-z.](https://doi.org/10.1186/s12864-024-10699-z)

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Authors' contributions

S.X. Wrote the manuscript text, prepared all fgures, and wrote the computer code. J.O. Brought the question to S.X., searched the literature, and edited the manuscript. All authors reviewed the manuscript.

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